

# Sex Modifies the *APOE*-Related Risk of Developing Alzheimer Disease

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**Objective:** The *APOE4* allele is the strongest genetic risk factor for sporadic Alzheimer disease (AD). Case-control studies suggest the *APOE4* link to AD is stronger in women. We examined the *APOE4*-by-sex interaction in conversion risk (from healthy aging to mild cognitive impairment (MCI)/AD or from MCI to AD) and cerebrospinal fluid (CSF) biomarker levels.

**Methods:** Cox proportional hazards analysis was used to compute hazard ratios (HRs) for an *APOE*-by-sex interaction on conversion in controls ( $n = 5,496$ ) and MCI patients ( $n = 2,588$ ). The interaction was also tested in CSF biomarker levels of 980 subjects from the Alzheimer's Disease Neuroimaging Initiative.

**Results:** Among controls, male and female carriers were more likely to convert to MCI/AD, but the effect was stronger in women (HR = 1.81 for women; HR = 1.27 for men; interaction:  $p = 0.011$ ). The interaction remained significant in a predefined subanalysis restricted to *APOE3/3* and *APOE3/4* genotypes. Among MCI patients, both male and female *APOE4* carriers were more likely to convert to AD (HR = 2.16 for women; HR = 1.64 for men); the interaction was not significant ( $p = 0.14$ ). In the subanalysis restricted to *APOE3/3* and *APOE3/4* genotypes, the interaction was significant ( $p = 0.02$ ; HR = 2.17 for women; HR = 1.51 for men). The *APOE4*-by-sex interaction on biomarker levels was significant for MCI patients for total tau and the tau-to- $A\beta$  ratio ( $p = 0.009$  and  $p = 0.02$ , respectively; more AD-like in women).

**Interpretation:** *APOE4* confers greater AD risk in women. Biomarker results suggest that increased *APOE*-related risk in women may be associated with tau pathology. These findings have important clinical implications and suggest novel research approaches into AD pathogenesis.

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Alzheimer disease (AD) is an increasingly prevalent, fatal neurodegenerative disease that has proven resistant thus far to all attempts to prevent it, forestall it, or slow its progression. The  $\epsilon 4$  allele of the apolipoprotein E gene (*APOE4*) is a potent genetic risk factor for sporadic and late onset familial AD.<sup>1</sup> The  $\epsilon 3$  allele (*APOE3*) is the most common *APOE* polymorphism in the general population and considered risk-neutral, whereas the  $\epsilon 2$  allele (*APOE2*) is the least common and is thought to reduce AD risk. Although estimates vary across studies and ethnic backgrounds, the *APOE4* allele is typically present in >50% of AD patients but is found only in about 15% of healthy older controls.<sup>2</sup> Basic science research has suggested several roles that the  $\epsilon 4$  isoform of apolipoprotein E (ApoE4) may play in

augmenting the development of AD. Cell culture and animal models have identified potential pathogenic mechanisms related to  $\beta$ -amyloid ( $A\beta$ ) clearance, tau hyperphosphorylation, and synaptic function, among others.<sup>3</sup>

In human studies, some but not all imaging biomarker studies have shown early AD-like findings in healthy older *APOE4* carriers.<sup>4–6</sup> Cerebrospinal fluid (CSF) biomarker studies are more consistent and tend to show reduced (more AD-like)  $A\beta$  levels, but normal tau levels, in healthy older *APOE4* carriers. Longitudinal studies of clinical decline from mild cognitive impairment (MCI) to AD are mixed, with some but not all suggesting that the *APOE4* allele increases the risk of conversion from MCI to AD.<sup>7</sup> The data on clinical

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conversion from healthy aging to MCI or AD are similarly mixed. To date, there have been 6 longitudinal studies examining the role of *APOE4* in the risk of converting from healthy aging to MCI or AD.<sup>8–13</sup> Of these studies, 4 found a significant effect of *APOE4* and 2 did not, even when combining *APOE4* heterozygotes and homozygotes. Thus, although the link between *APOE4* and AD is strong, many expected effects, like increasing the risk of conversion from healthy aging to MCI or from MCI to AD, have not been widely replicable.

A critical and commonly overlooked feature of the *APOE4* link to AD is that several case–control studies suggest it is far more pronounced in women. Shortly after the initial linkage studies, a prominent interaction between *APOE* and sex was reported.<sup>14</sup> The first large meta-analysis of *APOE4* studies confirmed the interaction and found that the effect was most prominent among subjects with 1 copy of the *APOE4* allele and 1 copy of the risk-neutral *APOE3* allele. Women with 1 *APOE4* allele had up to a 4-fold increased risk when compared to women homozygous for the *APOE3* allele. By contrast, men with 1 *APOE4* allele had little to no increase in risk.<sup>15</sup> This finding has been replicated and yet is rarely considered in clinical AD research, where male and female *APOE4* carriers are generally viewed as having equal risk.<sup>16,17</sup>

Although case–control studies of AD support an interaction between *APOE4* and sex, such studies are less conclusive than prospective cohort studies, particularly in diseases like AD with a long preclinical phase.<sup>18,19</sup> The interaction between *APOE4* and sex has not been established either in prospective cohorts of healthy older controls converting to MCI or AD or in prospective cohorts of MCI patients converting to AD. Most prospective studies examining the main effect of *APOE4* on incident MCI or AD have included sex as a covariate, but not explicitly tested for an *APOE4*-by-sex interaction. To our knowledge, only 1 prospective study has examined the effect of this interaction on clinical conversion, in this case from healthy aging to AD. Beydoun and colleagues reported a main effect of *APOE4* but no significant interaction with sex.<sup>8</sup> As acknowledged by the authors, this study, with 113 incident cases of all-cause dementia, may not have been adequately powered to detect a sex interaction.

In the current study, we hypothesized that the *APOE*-by-sex interaction would be evident in the risk of converting from healthy aging to MCI/AD and from MCI to AD and specifically that a single *APOE4* allele would confer greater risk of conversion in women than in men. To test this, we took advantage of a large, multi-

site, longitudinal aging and dementia database available through the National Alzheimer's Coordinating Center (NACC). In addition, to explore potential biochemical changes underlying these hypothesized effects, we also examined the *APOE*-by-sex interaction in CSF data from healthy older controls and MCI patients available in a second, multisite aging and dementia data set, the Alzheimer's Disease Neuroimaging Initiative (ADNI). In this case, we hypothesized that a single copy of the *APOE4* allele would result in more AD-like changes in women than in men (ie, lower  $A\beta$  levels, higher tau levels, higher phosphorylated-tau levels, and higher tau-to- $A\beta$  ratio).

## Subjects and Methods

### Assessing Conversion Risk in the NACC Data Set

The NACC curates data collected at 34 past and present AD centers across the United States. For this study, we used data from 11,654 nondemented subjects collected in the longitudinal Uniform Data Set,<sup>20–22</sup> with visits spanning from September 2005 to May 2013 (date of database access: June 12, 2013).

We restricted our analysis to subjects who were rated as healthy control (HC) or MCI at their initial assessment at study entry, who had an *APOE* genotype available, and who had a minimum of 1 follow-up visit at 12 months or later. These filter criteria led to a cohort of 8,084 subjects (HC: n = 5,496; MCI: n = 2,588; Table 1).

The clinical conversion risk was modeled using a Cox proportional hazards model. We performed the Cox regression analysis in 2 subgroups of the cohort: (1) HC only and (2) MCI only. For controls, clinical conversion was defined as the first detection of MCI or AD (using primary diagnoses of possible or probable AD based on the National Institute of Neurological and Communicative Diseases and Stroke–Alzheimer's Disease and Related Disorders Association criteria<sup>23</sup>), and the Cox model was used to model the hazard of developing MCI or AD among controls, whichever occurred first. For MCI subjects, conversion was defined as the first detection of AD (using primary diagnoses of possible or probable AD), and the Cox model was used to model the hazard of developing AD among MCI subjects. Any additional outcomes (such as reversions from MCI to HC or development of non-AD dementia) were treated as nonconversions. The subject age at the visit where clinical progression was detected served as the time of event; subjects who had not progressed at their last recorded visit were right-censored. In addition, to account for the period prior to inclusion in the cohort, a left-truncated design was used. Furthermore, we corrected the dependence between truncation and failure time (detected using Tsai's test<sup>24</sup>) by estimating time-dependent effects of the age at study entry in the Cox regression. This method is a generalization of the one proposed by Gail et al.<sup>25</sup> More precisely, we included age as a covariate in the Cox regression model and allowed the coefficient of age to be different within follow-up periods starting before and after

**TABLE 1. Population Statistics for HC and MCI Subjects in the National Alzheimer's Coordinating Center Data**

Characteristic	Total	$\epsilon 2/\epsilon 2$	$\epsilon 2/\epsilon 3$	$\epsilon 2/\epsilon 4$	$\epsilon 3/\epsilon 3$	$\epsilon 3/\epsilon 4$	$\epsilon 4/\epsilon 4$	$p$ , $APOE$	$p$ , Sex
HC									
Total No.	5,496	30	661	139	3,210	1,320	136	–	–
Female, No. (%)	3,652 (66.4)	19 (63.3)	456 (69.0)	96 (69.1)	2,117 (66.0)	872 (66.1)	92 (67.6)	0.71	–
Age, yr [IQR]	73.0 [66.6–79.8]	74.1 [68.1–81.8]	74.6 [67.1–81.8]	70.8 [65.3–77.8]	73.9 [67.4–80.4]	71.1 [65.5–78.1]	67.6 [62.1–73.3]	0.57	0.006 <sup>a</sup>
Visits, No. [IQR]	4 [3–6]	3.5 [3–4]	4 [3–6]	4 [3–5.5]	4 [3–6]	4 [3–6]	4 [2.8–6]	0.06	0.06
Follow-up, yr [IQR]	3.9 [2.2–5.5]	3.6 [2.2–4.3]	4.0 [2.3–5.7]	4.0 [2.6–5.2]	4.0 [2.2–5.5]	3.9 [2.1–5.4]	3.3 [2.1–5.3]	0.27	0.40
Converters, No. {%female}	964 {60.5}	2 {0}	111 {70.3}	29 {65.5}	521 {56.6}	268 {63.1}	33 {66.7}	0.08	<0.001 <sup>a</sup>
Education, yr [IQR]	16 [13–18]	16 [14–18]	16 [13–18]	16 [13.5–18]	16 [14–18]	16 [14–18]	16 [14–18]	0.63	<0.001 <sup>a</sup>
Non-Hispanic white, No.(%)	4,458 (81.1)	20 (66.7)	524 (79.3)	101 (72.7)	2,654 (82.7)	1,047 (79.3)	112 (82.4)	0.06	<0.001 <sup>a</sup>
MMSE [IQR]	29 [28–30]	29 [28–30]	29 [28–30]	29 [28–30]	29 [28–30]	29 [28–30]	29 [28–30]	0.68	<0.001 <sup>a</sup>
MCI									
Total No.	2,588	5	226	70	1,237	858	192	–	–
Female, No. (%)	1,275 (49.3)	3 (60.0)	117 (51.8)	35 (50.0)	602 (48.7)	425 (49.5)	93 (48.4)	0.63	–
Age, yr [IQR]	74.5 [68.5–80.2]	82.7 [78.2–84.2]	76.5 [70.0–83.2]	74.1 [67.3–82.3]	75.7 [68.8–81.6]	73.8 [68.6–79.0]	70.5 [65.9–74.5]	0.43	1.0
Visits, No. [IQR]	4 [2–5]	3 [2–3]	3 [3–5]	3 [2–5]	4 [3–5]	3 [2–5]	4 [3–5]	0.21	0.11
Follow-up, yr [IQR]	3.0 [2.0–4.7]	2.1 [2.1–3.5]	3.05 [2.03–4.68]	2.4 [1.6–4.4]	3.1 [2.0–4.8]	3.0 [1.8–4.4]	3.2 [2.1–4.9]	0.53	0.9
Converters, No. {%female}	867 {48.8}	0 {0}	50 {56.0}	23 {43.5}	324 {45.1}	364 {49.5}	106 {45.3}	0.04 <sup>a</sup>	0.73
Education, yr [IQR]	16 [12–18]	16 [12–16]	16 [12–18]	16 [12–18]	16 [12–18]	16 [12–18]	16 [14–18]	0.84	<0.001 <sup>a</sup>
Non-Hispanic white, No. (%)	2,071 (80.0)	2 (40)	176 (77.9)	58 (82.9)	967 (78.2)	704 (82.1)	164 (85.4)	0.05	<0.001 <sup>a</sup>
MMSE [IQR]	28 [26–29]	28 [28–29]	28 [26–29]	28 [26–29]	28 [26–29]	28 [26–29]	27 [25–29]	0.31	0.19

Columns show total subjects, subjects in each genotype, the p-value for the variable being different between  $\epsilon 2/\epsilon 3$  and  $\epsilon 3$  homozygotes, and the p-value for the characteristic being different between  $\epsilon 3$  homozygotes and  $\epsilon 3/\epsilon 4$  heterozygotes. P-values were computed using a logistic regression. Age refers to age at entry into the cohort. Number of visits refers to the number of visits recorded in the data-base. For continuous values, median and IQR are provided; for categorical data, the raw counts as well as percentage are given.

<sup>a</sup>Statistically significant at an uncorrected threshold of 0.05.

HC = healthy control; IQR = interquartile range; MCI = mild cognitive impairment; MMSE = Mini-Mental State Examination; – = not applicable.

the median age of the study sample (73.0 for controls and 74.5 for MCI; see Table 1). The data were analyzed across sexes with covariates for *APOE2* carrier status, *APOE4* carrier status, *APOE2* homozygosity, *APOE4* homozygosity, sex, *APOE2*-by-sex interaction, and *APOE4*-by-sex interaction. Due to the small sample size (see Table 1), *APOE2* homozygosity was not modeled in MCI subjects. Furthermore, the model was adjusted for years of education and Mini-Mental State Examination score at study entry, and stratified by race and Hispanic origin by grouping all non-Hispanic white subjects in one group and everyone else in a second group. In addition to the full regression model including the *APOE2*-by-sex and *APOE4*-by-sex variables, we assessed the risk of clinical conversion attributable to carrying the *APOE2* or *APOE4* allele for each sex separately in sex-stratified models. Hazard ratios (HRs) along with their 95% confidence intervals (CIs) are reported. To visualize conversion rates with left-truncated and right-censored data, we used the Breslow method to estimate the survival function of the conversion time for subjects entering the study at age 55 years.

Given our hypothesis, based on the Farrer et al meta-analysis,<sup>15</sup> that the *APOE4*-by-sex interaction would be strongest in  $\epsilon3/\epsilon4$  heterozygotes versus  $\epsilon3$  homozygotes, we also performed the Cox regression on a restricted subset of individuals having either the  $\epsilon3$  homozygote ( $\epsilon3/\epsilon3$ ) or the  $\epsilon3/\epsilon4$  heterozygote genotype.

### Effects on Spinal Fluid Biomarkers in the ADNI Database

We studied whether sex modulates the *APOE4* carrier effect on the most commonly used CSF biomarkers of AD. This analysis used data from the ADNI database ([www.loni.usc.edu/ADNI](http://www.loni.usc.edu/ADNI); date accessed: November 4, 2013). See Weiner et al<sup>26</sup> for an overview of the ADNI cohort. It should be noted that some ADNI subjects are likely also included in the NACC data set, although it is not currently possible to identify which subjects are in both data sets.

### CSF Biomarkers

In ADNI, a particular emphasis was placed on 4 established CSF biomarkers:  $A\beta$ , total tau, phosphorylated tau<sub>181p</sub> (p-tau), and the ratio of total tau to  $A\beta$ . Biomarkers were assessed at study entry and at follow-up visits for a subset of subjects (see Shaw et al<sup>5</sup> for details on biomarker acquisition).

CSF biomarkers were available for  $n = 1,094$  subjects at study entry. Again, the analysis was restricted to the HC ( $n = 272$ ) and MCI ( $n = 618$ ) subgroups. Table 2 lists sample sizes for each genotype and clinical category along with population demographics.

The *APOE4*-by-sex interaction on CSF biomarker levels at study entry was examined separately in HC and MCI subjects using an analysis of covariance (ANCOVA) adjusting for *APOE2* and *APOE4* carrier status, *APOE4* homozygosity, *APOE2*-by-sex interaction, sex, age, age-squared, years of education, and ADNI study phase (ANDI1 or ADNIGO/ADNI2). Due to the small sample sizes (see Table 2), *APOE2* homo-

zygosity was modeled neither in the HC nor in the MCI analyses. P-values for the 8 tests (ie, 4 CSF biomarkers for each HC and MCI) for an *APOE4*-by-sex interaction were corrected for multiple testing using the Holm–Bonferroni method. In an effort to examine potential  $A\beta$ -independent effects of *APOE4*, we report the total tau and p-tau analyses before and after adjusting for  $A\beta$  levels.

Again, given our hypothesis of a stronger *APOE4*-by-sex interaction in  $\epsilon3/\epsilon4$  heterozygotes versus  $\epsilon3$  homozygotes, we performed the ANCOVA on a restricted subset of individuals having either the  $\epsilon3$  homozygote ( $\epsilon3/\epsilon3$ ) or the  $\epsilon3/\epsilon4$  heterozygote.

## Results

### Clinical Conversion Risks

During the observation period, 959 healthy subjects (17.5%) converted to MCI or AD. The *APOE4*-by-sex interaction is significant ( $p = 0.011$ ) and driven by an increased *APOE4* effect on conversion in women. Among healthy older men, there is a marginally significant increase in conversion risk for *APOE4* carriers compared to noncarriers (HR = 1.27, 95% CI = 1.01–1.59,  $p = 0.045$ ). Among healthy older women, *APOE4* carriers show a highly significant 1.8-fold increase in risk (95% CI = 1.5–2.16,  $p = 2.5e-10$ ). The survival function plot (Fig 1) shows that noncarrier females are the group with the least risk in progressing from healthy control to MCI or AD (median conversion age (MCA) = 80.9 years, 95% CI = 79.7–83.0). *APOE4* carriers, regardless of sex, show the highest risk for clinical progression (MCA = 68.7, 95% CI = 67.6–70.7 and MCA = 71.6, 95% CI = 69.7–72.9 for males and females, respectively), whereas noncarrier males are at intermediate risk (MCA = 74.5, 95% CI = 70.2–77.6).

Among MCI subjects, the *APOE4* effect was significant in men and women, and the interaction was not significant ( $p = 0.14$ ; Table 3). Male *APOE4* carriers show an HR of 1.64 (95% CI = 1.33–2.02,  $p = 3.8e-6$ ), whereas female *APOE4* carriers exhibit an HR of 2.16 (95% CI = 1.74–2.69,  $p = 2.9e-12$ ; Fig 2).

The *APOE2*-by-sex interaction is significant only in healthy controls ( $p = 0.045$ ). Here, men carrying the *APOE2* allele show a nonsignificant decrease in risk (HR = 0.74, 95% CI = 0.53–1.01,  $p = 0.06$ ), whereas female *APOE2* carriers exhibit a nonsignificant increase in risk (HR = 1.12, 95% CI = 0.90–1.40,  $p = 0.32$ ).

Restriction of the analysis to  $\epsilon3/\epsilon3$  and  $\epsilon3/\epsilon4$  subjects confirms the *APOE4*-by-sex interaction in healthy controls ( $p = 0.02$ ; Fig 3) and reveals a significant *APOE4*-by-sex interaction among MCI subjects ( $p = 0.02$ ; Fig 4). Among controls, a single copy of the *APOE4* allele significantly increases the conversion risk in women (HR = 1.8, 95% CI = 1.48–2.19,  $p = 3.5e-9$ )

**TABLE 2. Population Statistics for HC and MCI Subjects in the Alzheimer's Disease Neuroimaging Initiative Database for Whom Cerebrospinal Fluid Biomarker Data Were Available**

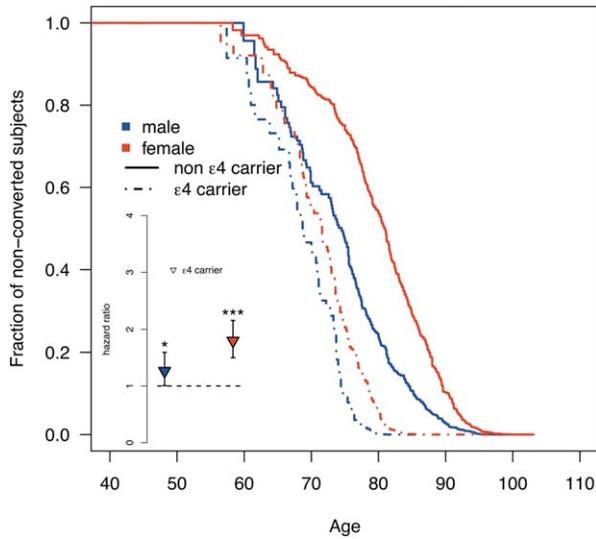
Characteristic	Total	ε2/ε2	ε2/ε3	ε2/ε4	ε3/ε3	ε3/ε4	ε4/ε4	p, APOE	p, Sex
<b>HC</b>									
Total No.	272	0	39	2	162	61	8	-	-
Age, yr [IQR]	73.7 [70.8–78.1]	-	72.4 [70.6–76.3]	72.3 [71.3–73.2]	74.0 [71.0–78.5]	73.7 [68.2–77.5]	76.8 [69.0–83.3]	0.42	0.11
Female, No. (%)	135 (49.8)	-	22 (56.4)	1 (50)	79 (48.8)	30 (49.2)	3 (38)	0.37	-
Education, yr [IQR]	16 [14–18]	-	16 [13.5–18]	14.5 [13.3–17.8]	16 [14–18]	16 [14–18]	17 [16–18.5]	0.29	<0.001 <sup>a</sup>
MMSE [IQR]	29 [29–30]	-	29 [28–30]	27.5 [26.3–28.8]	29 [29–30]	28 [29–30]	29.5 [29–30]	0.31	0.04 <sup>a</sup>
<b>MCI</b>									
Total No.	618	1	36	10	274	230	67	-	-
Age, yr [IQR]	72.8 [67.3–77.6]	77.6	73.2 [69.8–78.3]	70.0 [64.3–74.6]	74.0 [67.9–79.7]	72.7 [67.5–76.7]	70.0 [64.7–73.9]	0.48	<0.001 <sup>a</sup>
Female, No. (%)	249 (40.5)	1 (100)	13 (36.1)	5 (50)	110 (40.3)	94 (41.0)	26 (39.4)	0.18	-
Education, yr [IQR]	16 [14–18]	17	16 [13.8–18]	17.5 [14.3–18]	16 [14–18]	16 [14–18]	16 [14–18]	0.73	<0.001 <sup>a</sup>
MMSE [IQR]	28 [26–29]	29	28 [27–29]	28 [26.3–29.8]	28 [27–29]	28 [26–29]	28 [26–30]	0.43	0.14

Columns show total subjects and subjects for each APOE genotype. Age refers to age at entry into the cohort. For continuous values, median and IQR are provided; for categorical data, the raw counts as well as percentage are given. The last 2 columns contain the p-value for the variables being different between APOE genotypes and sex, respectively. P-values were computed using logistic regression.

<sup>a</sup>Statistically significant at an uncorrected threshold of 0.05.

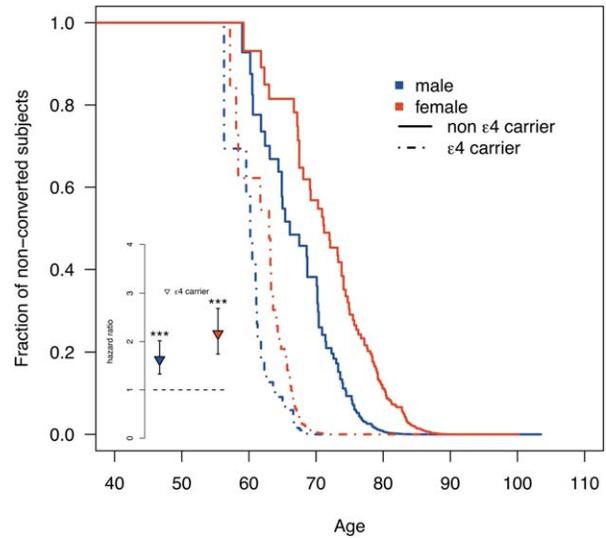
HC = healthy control; IQR = interquartile range; MCI = mild cognitive impairment; MMSE = Mini-Mental State Examination; - = not applicable.

but not in men (HR = 1.23, 95% CI = 0.96–1.57,  $p = 0.09$ ). Among MCI patients, a single copy of the *APOE4* allele significantly increases conversion risk in



**FIGURE 1:** The *APOE4* carrier status increases the risk of clinical decline in healthy older women more than in men. The main figure shows the survival function plot for conversions from healthy control (HC) to either mild cognitive impairment (MCI) or possible/probable Alzheimer disease (AD) for subjects entering the study at age 55 years based on left-truncated and right-censored data. In contrast to the Cox regression, no stratification or covariate adjustment was applied. The inset depicts the hazard ratio for converting from HC to MCI or AD computed separately for each sex using a Cox regression model (significant effects at  $*p \leq 0.05$  and  $***p < 0.001$ ). Blue and red refer to males and females, respectively.

men (HR = 1.51, 95% CI = 1.22–1.88,  $p = 2e-4$ ) and women (HR = 2.17, 95% CI = 1.72–2.72,  $p = 3.6e-11$ ) but, as indicated by the significant interaction, to a significantly greater degree in women.

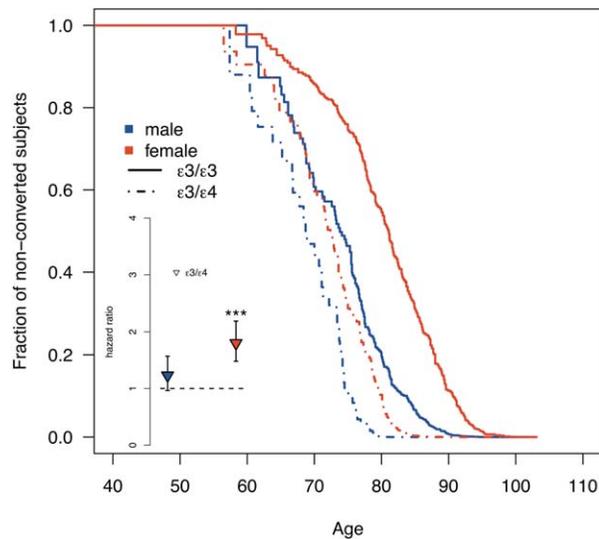


**FIGURE 2:** The *APOE4* carrier status increases the risk of clinical decline in both women and men with mild cognitive impairment (MCI). The main figure shows the survival function plot for conversions from MCI to possible/probable Alzheimer disease (AD) for subjects entering the study at age 55 years based on left-truncated and right-censored data. In contrast to the Cox regression, no stratification or covariate adjustment was applied. The inset depicts the hazard ratio for converting from MCI to AD computed separately for each sex using a Cox regression model ( $***p < 0.001$ ). Blue and red refer to males and females, respectively.

**TABLE 3. Cox Proportional HRs**

Covariate	HC			MCI		
	Cox HR	95% CI	$p$	Cox HR	95% CI	$p$
Edu	0.98	0.96–1.00	0.12	1.04	1.01–1.06	0.00115
MMSE	0.85	0.82–0.89	2.2e-13	0.83	0.81–0.85	<2.2e-16
Sex	0.69	0.58–0.82	3.2e-05	0.98	0.79–1.22	0.86
<i>APOE2</i>	0.75	0.55–1.02	0.069	0.83	0.59–1.19	0.31
<i>APOE2</i> hom	0.48	0.12–1.94	0.30	–	–	–
<i>APOE4</i>	1.25	1.00–1.57	0.048	1.70	1.39–2.07	1.7e-07
<i>APOE4</i> hom	1.59	1.11–2.29	0.012	1.39	1.11–1.74	0.004
<i>APOE2</i> ×sex	1.48	1.01–2.17	0.045	0.99	0.61–1.62	0.97
<i>APOE4</i> ×sex	1.44	1.09–1.91	0.011	1.23	0.94–1.63	0.14

Rows correspond to the covariates in the Cox regression models. Columns correspond to the Cox proportional hazards ratio, its 95% CI, and the corresponding  $p$ -value for the analysis including all subjects for HC and MCI subjects. *APOE2* = *APOE2* carrier status; *APOE2*hom = *APOE2* homozygosity; *APOE4* = *APOE4* carrier status; *APOE4*hom = *APOE4* homozygosity; CI = confidence interval; Edu = years of education; HC = healthy control; HR = hazard ratio; MCI = mild cognitive impairment; MMSE = Mini-Mental State Examination; – = not applicable.



**FIGURE 3:** A single *APOE4* allele increases the risk of clinical decline in healthy older women, but not men. The main figure shows the survival function plot for conversions from healthy control (HC) to either mild cognitive impairment (MCI) or possible/probable Alzheimer disease (AD) for subjects with either the  $\epsilon 3/\epsilon 3$  or the  $\epsilon 3/\epsilon 4$  genotype entering the study at age 55 years based on left-truncated and right-censored data. In contrast to the Cox regression, no stratification or covariate adjustment was applied. The inset depicts the hazard ratio for converting from HC to MCI or AD computed separately for each sex using a Cox regression model (\*\*\*)significant effect at  $p < 0.001$ ). Blue and red refer to males and females, respectively.

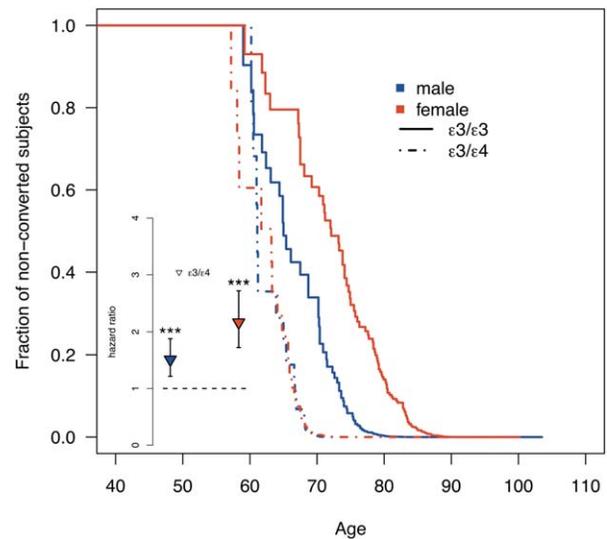
### CSF Biomarkers

The *APOE4*-by-sex interaction showed only a nominally significant effect ( $p = 0.02$ ,  $p_{\text{cor}} = 0.12$ ) for CSF  $A\beta$  levels in healthy controls (more AD-like in men; Fig 5A); there was no evidence for an interaction by sex on the other 3 biomarkers among controls (Table 4). Among MCI subjects, we detected a significant *APOE4*-by-sex interaction on CSF tau levels ( $p_{\text{cor}} = 0.009$ ) and on the tau- $A\beta$ -ratio ( $p_{\text{cor}} = 0.02$ ) as well as a nominally significant effect on CSF p-tau ( $p = 0.025$ ,  $p_{\text{cor}} = 0.13$ ) with a more AD-like pattern in female *APOE4* carriers (see Fig 5C, D). Results for tau and p-tau remained qualitatively unchanged after adjusting for  $A\beta$  (see Table 4).

The restricted analysis on  $\epsilon 3$  homozygous and  $\epsilon 3/\epsilon 4$  heterozygous subjects confirmed the same *APOE4*-by-sex interactions (data not shown).

### Discussion

We have demonstrated, in a large, longitudinal sample, that the risk of clinical conversion conferred by the *APOE4* allele is significantly greater in women than in men. The interaction was present in the conversion from healthy aging to MCI/AD and in the conversion from MCI to AD. Among healthy controls, the interaction



**FIGURE 4:** A single *APOE4* allele increases the risk of clinical decline in women with mild cognitive impairment (MCI) more than in men. The main figure shows the survival function plot for conversions from MCI to possible/probable Alzheimer disease (AD) for subjects with either the  $\epsilon 3/\epsilon 3$  or the  $\epsilon 3/\epsilon 4$  genotype entering the study at age 55 years based on left-truncated and right-censored data. In contrast to the Cox regression, no stratification or covariate adjustment was applied. The inset depicts the hazard ratio for converting from MCI to AD computed separately for each sex using a Cox regression model (\*\*\*)significant effect at  $p < 0.001$ ). Blue and red refer to males and females, respectively.

was detectable both in the full analysis (including all genotypes) and in the predefined subanalysis restricted to the 2 most common genotypes ( $\epsilon 3$  homozygotes vs  $\epsilon 3/\epsilon 4$  heterozygotes, accounting for 82% of controls). In this sample, healthy older male *APOE4* carriers were at a marginally significant increased risk of converting to MCI or AD when compared with men who did not carry the *APOE4* allele. By contrast, healthy older female *APOE4* carriers were almost twice as likely to develop MCI or AD when compared to female noncarriers. In the subanalysis, the *APOE4* effect remained significant in women but was no longer significant in men. Among all MCI subjects, *APOE4* carriers of both sexes had an increased risk of conversion to AD. In the subanalysis ( $\epsilon 3$  homozygotes vs  $\epsilon 3/\epsilon 4$  heterozygotes, accounting for 81% of MCI subjects), the *APOE4* effect was significant in men and women but was significantly stronger in women (*APOE4*-by-sex interaction:  $p = 0.02$ ). These prospective findings on clinical conversion support earlier case-control analyses demonstrating that women with a single *APOE4* allele were at increased risk of developing AD compared to women who were homozygous for the *APOE3* allele, whereas men with a single *APOE4* allele were not at increased risk when compared to men who were *APOE3* homozygotes.<sup>15-17</sup> Although the *APOE2*

TABLE 4. P-values from the Analysis of Variance on CSF Biomarker Levels

Biomarker	Sex	<i>APOE2</i>	<i>APOE2</i> ×sex	<i>APOE4</i>	<i>APOE4</i> hom	<i>APOE4</i> ×sex	<i>p</i> <sub>cor</sub>
HC							
Aβ	0.07	0.002	0.23	0.05	0.008	0.02	0.12
Tau	0.24	0.52	0.97	0.018	0.07	0.62	1.0
p-Tau	0.90	0.44	0.54	0.001	0.30	0.99	1.0
Tau/Aβ	0.41	0.41	0.86	0.002	0.88	0.83	1.0
Tau <sup>a</sup>	0.38	0.94	0.86	0.05	0.02	0.37	–
p-Tau <sup>a</sup>	0.76	0.13	0.36	0.005	0.09	0.58	–
MCI							
Aβ	0.71	0.33	0.13	2.3e-13	1.1e-07	0.21	0.84
Tau	0.70	0.28	0.45	4.6e-13	0.34	0.001	0.009 <sup>b</sup>
p-Tau	0.52	0.54	0.94	2.2e-10	0.24	0.025	0.13
Tau/Aβ	0.71	0.27	0.38	<2e-14	0.005	0.003	0.02 <sup>b</sup>
Tau <sup>a</sup>	0.56	0.39	0.74	3.7e-07	0.49	0.003	–
p-Tau <sup>a</sup>	0.58	0.80	0.48	2.7e-04	0.37	0.06	–

Columns correspond to the main covariates of interest: sex, *APOE2* carrier status, *APOE2*-by-sex interaction, *APOE4* carrier status, *APOE4* homozygosity, and *APOE4*-by-sex interaction. The last column reports the corrected p-value for the *APOE4*-by-sex interaction for the 8 main tests. Rows correspond to different CSF biomarkers: Aβ, Tau, p-Tau, and the Tau/Aβ ratio.

<sup>a</sup>Results for tau and p-tau were corrected for Aβ levels.

<sup>b</sup>Statistically significant at a corrected threshold of 0.05.

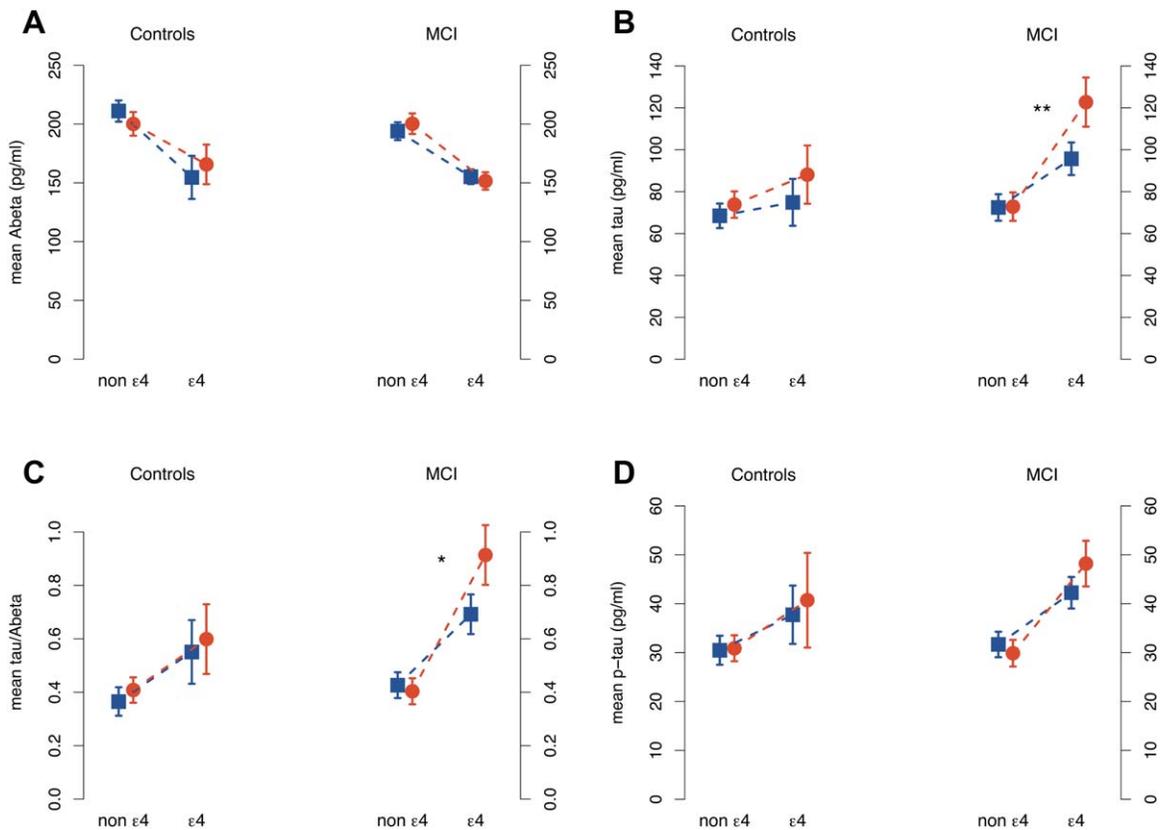
Aβ = β-amyloid; *APOE4*hom = *APOE4* homozygosity; CSF = cerebrospinal fluid; HC = healthy control; MCI = mild cognitive impairment; – = not applicable; p-Tau = phosphorylated tau.

findings should be considered preliminary, owing to the smaller sample sizes, here too we detected a significant interaction with sex in which the *APOE2* allele trended toward being protective in male, but not female, controls.

The CSF biomarker results reported here suggest that the increased risk of AD in female *APOE4* carriers occurs downstream of Aβ pathology. Aβ pathology is believed to occur early during disease pathogenesis, before the appearance of tau-related changes reflective of neuronal injury. The effect of carrying an *APOE4* allele on lowering Aβ levels was quite pronounced in both healthy older men and women. Total tau and p-tau levels showed a main effect of *APOE4* but did not show any *APOE*-by-sex interactions in the healthy older controls. Among MCI patients, however, *APOE4* increased total tau levels significantly more in women than in men, even after controlling for Aβ levels. Similarly, among MCI patients, *APOE4* increased the ratio of total tau to Aβ significantly more in women than in men, despite similar Aβ levels. Although p-tau levels showed a nominally significant trend in the same direction (more AD-like in female *APOE4* carriers), this did

not survive Holm–Bonferonni correction for multiple comparisons.

One possible explanation for the increased effect of *APOE4* on the tau biomarker in women is that amyloid changes occur earlier in women than in men. This possibility is less likely given that our analyses adjusted for linear and quadratic effects of age on amyloid, but subsequent longitudinal studies can assess whether there is an earlier start to amyloid pathology in women with the *APOE4* allele. An alternative explanation is that for an equivalent amount and duration of amyloid pathology, the *APOE4* allele results in more tau-related pathology in women compared to men. *APOE4* may initially change Aβ processing in a manner that is roughly equivalent for both sexes but then triggers a more robust acceleration of tau pathology in women. This explanation also appears unlikely, because even after adjusting for the effect of amyloid on tau, the *APOE*-by-sex interaction is still significant. An Aβ-independent effect of *APOE* on CSF tau has also previously been shown by Cruchaga et al, where polymorphisms in the *APOE* locus were strongly associated with CSF tau levels even after an adjustment for Aβ levels.<sup>27</sup>



**FIGURE 5: Sex modifies the APOE4 effect on cerebrospinal fluid (CSF) biomarker levels.** CSF biomarker levels in healthy control (HC) and mild cognitive impairment (MCI) subjects are shown with 95% confidence intervals. Depicted CSF biomarker levels were adjusted for age, age-squared, years of education, Mini-Mental State Examination score, and Alzheimer's Disease Neuroimaging Initiative study phase. Blue squares and red circles correspond to men and women, respectively. Dashed lines highlight the change in CSF levels between  $\epsilon 3$  homozygotes and  $\epsilon 4$  heterozygotes. P-values for the APOE by sex effect were computed using an analysis of covariance. \* $p < 0.05$  (corrected for multiple comparisons), \*\* $p < 0.01$  (corrected). A–D correspond to  $\beta$ -amyloid, total tau, tau-to-A $\beta$ -ratio, and p-tau, respectively.

Two caveats should be considered when interpreting these results. First, our results may generalize imperfectly, as neither the NACC nor the ADNI data are population-based. For example, a recruitment bias could account for the unexpected finding that the risk of conversion from healthy aging to MCI or AD is less for female than male APOE3 homozygotes. This novel finding, although not predicted by our a priori hypothesis, is nonetheless worth pursuing in a population-based study. Second, the Cox model assumes that dropout and censoring are unrelated to conversion. Impaired subjects may have been more likely to drop out, and there may have been differential effects by sex.

Despite compelling evidence from a large meta-analysis of case–controls studies, the field of AD research has largely overlooked this potent interaction between APOE and sex.<sup>15</sup> This may have been due in part to the lack of any previous, prospective cohort studies supporting this interaction effect on clinical conversion.<sup>8</sup> It seems likely that a number of inconsistent findings related to APOE,

some of which are outlined in the introduction, are a result of investigators overlooking the APOE-by-sex interaction. We hope that the current findings will alert the field to this interaction and the important clinical and scientific implications it carries. From a clinical perspective, these results require careful reexamination of how we should interpret the finding of a single APOE4 allele in men. This bears importantly on encounters with individual patients in terms of diagnostics, prognostics, and genetic counseling. In regard to clinical trials, appreciating the APOE-by-sex interaction should allow for more refined genotype stratification when, for example, estimating conversion risk in preventative trials.<sup>28</sup> Furthermore, several drug trials have suggested that efficacy and side effect profiles may differ between APOE4 carriers and noncarriers, although these studies have not assessed the sex interaction.<sup>29</sup> From a scientific standpoint, these findings should motivate investigations into the potential mechanisms of the APOE-by-sex interaction. Explicitly modeling this interaction, both in human studies and in animal model studies, as is now only

occasionally done, has the potential to yield new insights into the strongest genetic risk factor for late onset AD.<sup>30–32</sup>

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## Authorship

M.D.G. conceived the study. M.D.G. and V.W.H. conducted literature search. A.A. collected data and prepared figures. A.A. and L.T. analyzed data. M.D.G. and A.A. wrote first draft. All authors contributed in analyzing the data and writing the final manuscript.

## Potential Conflicts of Interest

Nothing to report.

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